

## Introduction to mathematical Statistics

### Cramer-Rao lower bound

Unbiased Estimator of  $\theta$ , say  $\hat{\theta}_1, \hat{\theta}_2, \hat{\theta}_3, \dots$

It could be really difficult for us to compare  $Var(\hat{\theta}_i)$  when there are many of them.

### Theorem. Cramer-Rao Lower Bound

Let  $Y_1, Y_2, \dots, Y_n$  be a random sample from a population with p.d.f.  $f(y, \theta)$ . Let  $\hat{\theta} = h(y_1, y_2, \dots, y_n)$  be an unbiased estimator of  $\theta$ .

Given some regularity conditions (continuous differentiable etc.) and the domain of  $f(y_i, \theta)$  does not depend on  $\theta$ .

Then, we have

$$\text{var}(\hat{\theta}) \geq \frac{1}{n \cdot E \left[ \left[ \frac{\partial (\ln f(\theta))}{\partial \theta} \right]^2 \right]} = \frac{1}{-n \cdot E \left[ \frac{\partial^2 (\ln f(\theta))}{\partial \theta^2} \right]}$$

### **Theorem. Properties of the MLE**

Let  $Y_i \stackrel{i.i.d.}{\sim} f(y, \theta)$ ,  $i = 1, 2, \dots, n$

Let  $\hat{\theta}$  be the MLE of  $\theta$

Then,  $\hat{\theta} \xrightarrow{n \rightarrow \infty} N\left(\theta, \frac{1}{n \cdot E \left[ \left[ \frac{\partial \ln(f(\theta))}{\partial \theta} \right]^2 \right]}\right)$

The MLE is asymptotically unbiased and its asymptotic variance : C-R lower bound

**Example 1.** Let  $Y_1, Y_2, \dots, Y_n \stackrel{i.i.d.}{\sim} \text{Bernoulli}(p)$

1. MLE of  $p$  ?
2. What are the mean and variance of the MLE of  $p$  ?
3. What is the Cramer-Rao lower bound for an unbiased estimator of  $p$  ?

**Solution.**  $P(Y = y) = f(y; p) = p^y(1-p)^{1-y}$  ;  $y=0,1$

$$\begin{aligned} 1. \quad L &= \prod_{i=1}^n f(y_i; p) \\ &= \prod_{i=1}^n p^{y_i} (1-p)^{1-y_i} \\ &= p^{\sum y_i} (1-p)^{n-\sum y_i} \end{aligned}$$

$$l = \ln L = (\sum y_i) \ln p + (n - \sum y_i) \ln(1-p)$$

$$\frac{\partial l}{\partial p} = \frac{\sum y_i}{p} - \frac{n - \sum y_i}{1-p} = 0$$

$$\hat{p} = \frac{\sum_{i=1}^n Y_i}{n} : \text{MLE}$$

$$2. \quad E(\hat{p}) = p$$

$$\text{var}(\hat{p}) = \frac{p(1-p)}{n}$$

$$3. \quad \ln f(y, p) = y \ln p + (1-y) \ln(1-p)$$

$$\frac{\partial \ln f(y, p)}{\partial p} = \frac{y}{p} - \frac{1-y}{1-p}$$

$$\frac{\partial^2 \ln f(y, p)}{\partial p^2} = -\frac{y}{p^2} - \frac{1-y}{(1-p)^2}$$

$$E \left[ -\frac{Y}{p^2} - \frac{1-Y}{(1-p)^2} \right] = -\frac{p}{p^2} - \frac{1-p}{(1-p)^2} = -\frac{1}{p(1-p)}$$

C-R lower bound

$$\text{var}(\hat{p}) \geq \frac{1}{-nE\left[\frac{\partial^2 \ln f}{\partial p^2}\right]} = \frac{p(1-p)}{n}$$

The MLE of  $p$  is unbiased and its variance = C-R lower bound.

**Definition. Efficient Estimator**

If  $\hat{\delta}$  is an unbiased estimator of  $\delta$  and its variance = C-R lower bound, then  $\hat{\delta}$  is an efficient estimator of  $\delta$ .

**Definition. Best Estimator**

If  $\hat{\delta}$  is an unbiased estimator of  $\delta$  and  $\text{var}(\hat{\delta}) \leq \text{var}(\tilde{\delta})$  for all unbiased estimator  $\tilde{\delta}$ , then  $\hat{\delta}$  is a best estimator for  $\delta$ .

Always true  
Efficient Estimator  $\stackrel{\text{Always true}}{=}$  Best Estimator  
May not be true

**Example 2.** If  $Y_1, Y_2, \dots, Y_n$  is a random sample from  $f(y, \theta) = \frac{2y}{\theta^2}$ ,  $2 < y < \theta$ , then  $\hat{\theta} = \frac{3}{2}\bar{Y}$  is an unbiased estimator for  $\theta$ .

Compute 1.  $\text{Var}(\hat{\theta})$  and 2. C-R lower bound for  $f_y(y; \theta)$

**Solution.**

$$1. \text{Var}(\hat{\theta}) = \text{Var}\left(\frac{3}{2}\bar{Y}\right) = \frac{9}{4}\left(\frac{1}{n}\sum_{i=1}^n Y_i\right) = \frac{9}{4n^2}\sum_{i=1}^n \text{Var}(Y_i)$$

$$\text{Var}(Y_i) = E(Y_i^2) - [E(Y_i)]^2 = \int_0^\theta y^2 \frac{2y}{\theta^2} dy - \left[\int_0^\theta y^2 \frac{2y}{\theta^2} dy\right]^2 = \frac{\theta^2}{18}$$

$$\text{Therefore, } \text{Var}(\hat{\theta}) = \frac{9}{4n^2} \frac{n\theta^2}{18} = \frac{\theta^2}{8n}$$

2. C-R lower bound

$$\ln f_Y(y, \theta) = \ln\left(\frac{2y}{\theta^2}\right) = \ln 2y - 2 \ln \theta$$

$$\frac{\partial \ln f_Y(y, \theta)}{\partial \theta} = -\frac{2}{\theta}$$

$$E\left[\left(\frac{\partial \ln f_Y(y, \theta)}{\partial \theta}\right)^2\right] = E\left(\frac{4}{\theta^2}\right) = \int_0^\theta \frac{4}{\theta^2} \frac{2y}{\theta^2} dy = \frac{4}{\theta^2}$$

$$\therefore \frac{1}{nE\left[\left(\frac{\partial \ln f_Y(y, \theta)}{\partial \theta}\right)^2\right]} = \frac{\theta^2}{4n}$$

The value of 1 is less than the value of 2. But, it is **NOT** a contradiction to the theorem. Because  $0 < y < \theta$ , the domain depends on  $\theta$ . Thus the C-R Theorem doesn't hold for this problem.