

Example 2: A mixed model with count data

The following example is from SAS System for Mixed Models. The experiment is a split-plot with different management methods as the whole-plot treatment factor and different seed mixes as the split-plot treatment factor. The whole-plots were arranged in randomized complete blocks. The measurement was the number of plants of a given species per experimental unit. There are 7 management types (TRT), 4 blocks (BLK), 4 types of seed mix (MIX) and the response is the number of plants (COUNT). The data are located in the splitplotcount.dat dataset. Because it is count data, we will assume a Poisson distribution. The model that we will be fitting is

$$\log(\mu_{ijk}) = a + r_i + \tau_j + (r\tau)_{ij} + \delta_k + (\tau\delta)_{jk}$$

where

μ_{ijk} is the conditional mean count given the random effects
a is the intercept
 r_i is the BLK effect
 τ_j is the TRT effect
 $(r\tau)_{ij}$ is the BLK*TRT (whole-plot error) effect
 δ_k is the MIX effect
 $(\tau\delta)_{jk}$ is the TRT*MIX interaction

$$r_i \sim N(0, \sigma_r^2)$$

$$(r\tau)_{ij} \sim N(0, \sigma_{RT}^2)$$

The following SAS code reads in the data and prints out the first 19 observations

```
options ps=66 ls=80;
filename sp 'f:\kathy mixed model\splitplotcount.dat';
data spd;
infile sp;
input trt blk mix count;
run;

data select;
set spd;
if _n_ < 20;
proc print;
run;
```

The following is the printout of the first observations. There are 112 observations (7*4*4).

Obs	trt	blk	mix	count
-----	-----	-----	-----	-------

1	1	1	1	24
2	1	1	2	12
3	1	1	3	8
4	1	1	4	13
5	1	2	1	9
6	1	2	2	9
7	1	2	3	9
8	1	2	4	18
9	1	3	1	12
10	1	3	2	8
11	1	3	3	44
12	1	3	4	0
13	1	4	1	8
14	1	4	2	12
15	1	4	3	25
16	1	4	4	0
17	2	1	1	11
18	2	1	2	32
19	2	1	3	12

The GLIMMIX procedure is used for this analysis.

```
proc glimmix data=spd;  
class trt blk mix;  
model count=trt mix trt*mix/dist=poisson link=log ddfm=satterth;  
lsmeans trt mix/diff;  
random blk blk*trt;  
run;
```

The class statement specifies the classification effects in the model. Because we are using the Poisson distribution, the model dependent variable (left hand side) is the count information. The independent variables on the right side of the equation include the fixed effects, which are TRT, MIX, and the TRT*MIX interaction effect. The random statement specifies the random effects (BLK, and whole plot error BLK*TRT) that are included in the model.

The first information produced by this program is the model information.

Model Information

Data Set	WORK.SPD
Response Variable	count
Response Distribution	Poisson
Link Function	Log
Variance Function	Default
Variance Matrix	Not blocked
Estimation Technique	Residual PL
Degrees of Freedom Method	Satterthwaite

The model information is correct for the model specified. The next output is the class level information and the dimension information for the design matrices (X for fixed effects and Z for random effects). The number of covariance parameters is 2 (BLK and BLK*TRT). The columns in X are 40, which is 1 for the intercept, + 7 for TRT, +4 for MIX, + 7*4=28 for TRT*MIX. The columns in Z are 32, which is 4 for the BLK, +4*7=28 for BLK*TRT.

Class Level Information

Class	Levels	Values
trt	7	1 2 3 4 5 6 7
blk	4	1 2 3 4
mix	4	1 2 3 4

Number of Observations Read	112
Number of Observations Used	112

Dimensions

G-side Cov. Parameters	2
Columns in X	40
Columns in Z	32
Subjects (Blocks in V)	1
Max Obs per Subject	112

The next information presented is the information on the optimization. Because we cannot get a closed form solution for the variance components, we need to iteratively solve the equations until convergence is reached. The default optimization technique used in PROC GLIMMIX is the Dual Quasi-Newton. We are trying to estimate two variance components BLK and BLK*TRT which are in the G matrix. This analysis took five iterations to reach convergence. The output also has a message that the Estimated G matrix is not positive definite. This indicates that one of the variance components that was being estimated had a negative estimate. SAS automatically sets negative estimates to 0.

Optimization Information

Optimization Technique	Dual Quasi-Newton
Parameters in Optimization	2

Lower Boundaries 2
 Upper Boundaries 0
 Fixed Effects Profiled
 Starting From Data

Iteration History

Iteration	Restarts	Subiterations	Objective Function	Change	Max Gradient
0	0	4	472.45598993	2.00000000	44.26195
1	0	4	594.57559733	0.59355256	50.0345
2	0	2	624.66008319	0.06096869	50.37533
3	0	2	626.86889196	0.00056067	50.36349
4	0	1	626.88648763	0.00000153	50.36334
5	0	0	626.88649572	0.00000000	50.36334

Convergence criterion (PCONV=1.11022E-8) satisfied.

Estimated G matrix is not positive definite.

The next information presented is the goodness of fit. The only fit statistics that we will focus on is the Pearson Chi-Square/DF. This gives us an estimate of the overdispersion parameter. This parameter is telling us that there is quite a bit of overdispersion and that the statistical tests will be liberal.

Fit Statistics

-2 Res Log Pseudo-Likelihood	626.89
Pseudo-AIC (smaller is better)	628.89
Pseudo-AICC (smaller is better)	628.94
Pseudo-BIC (smaller is better)	628.27
Pseudo-CAIC (smaller is better)	629.27
Pseudo-HQIC (smaller is better)	627.54
Pearson Chi-Square	634.08
Pearson Chi-Square / DF	7.55

The covariance parameter estimates shows us that the variance estimate for BLK was the parameter that was set to zero. Remember that the TRT effect is tested over the TRT*BLK variance estimate.

Covariance Parameter Estimates

Cov Parm	Estimate	Standard Error
blk	5.93E-19	.
trt*blk	0.1652	0.05596

The Type III tests for the fixed effects are presented next. Note that the denominator degrees of freedom are 1, which does not make sense. This is because we asked for the Satterthwaite degrees of freedom which are based on the variance being fixed. SAS doesn't know how to handle the degrees of freedom calculation when you are not estimating overdispersion, so the

degrees of freedom are nonsense. If we ran the same model without the ddfm=satterth, the denominator degrees of freedom for MIX and TRT*MIX become 63 and the Pr > F for both of these become <.0001.

Type III Tests of Fixed Effects

Effect	Num DF	Den DF	F Value	Pr > F
trt	6	20.26	3.64	0.0130
mix	3	1	14.82	0.1882
trt*mix	18	1	10.48	0.2391

There is a significant treatment effect. Even though both MIX and TRT*MIX have large F values, neither of them are statistically significant. This is because the probabilities have taken into account the overdispersion.

The least-squares means are presented next.

trt Least Squares Means

trt	Estimate	Standard Error	DF	t Value	Pr > t
1	2.4973	0.2164	22.41	11.54	<.0001
2	2.8825	0.2129	20.91	13.54	<.0001
3	2.8279	0.2125	20.83	13.31	<.0001
4	3.0979	0.2106	20.08	14.71	<.0001
5	3.4404	0.2084	19.27	16.51	<.0001
6	3.2579	0.2100	19.85	15.51	<.0001
7	3.6928	0.2073	18.88	17.81	<.0001

Differences of trt Least Squares Means

trt	_trt	Estimate	Standard Error	DF	t Value	Pr > t
1	2	-0.3852	0.3036	21.65	-1.27	0.2180
1	3	-0.3307	0.3033	21.61	-1.09	0.2877
1	4	-0.6006	0.3020	21.23	-1.99	0.0597
1	5	-0.9431	0.3004	20.81	-3.14	0.0050
1	6	-0.7606	0.3016	21.11	-2.52	0.0198
1	7	-1.1955	0.2997	20.61	-3.99	0.0007
2	3	0.05451	0.3008	20.87	0.18	0.8580
2	4	-0.2155	0.2995	20.49	-0.72	0.4800
2	5	-0.5579	0.2979	20.08	-1.87	0.0757
2	6	-0.3755	0.2991	20.38	-1.26	0.2235
2	7	-0.8104	0.2972	19.88	-2.73	0.0130
3	4	-0.2700	0.2992	20.45	-0.90	0.3774
3	5	-0.6124	0.2977	20.04	-2.06	0.0529
3	6	-0.4300	0.2988	20.34	-1.44	0.1654
3	7	-0.8649	0.2969	19.84	-2.91	0.0087
4	5	-0.3425	0.2963	19.67	-1.16	0.2615
4	6	-0.1600	0.2974	19.97	-0.54	0.5966
4	7	-0.5949	0.2955	19.47	-2.01	0.0581
5	6	0.1825	0.2959	19.56	0.62	0.5445
5	7	-0.2524	0.2940	19.07	-0.86	0.4012
6	7	-0.4349	0.2951	19.36	-1.47	0.1567

There are several pairs of treatments that are different from each other. Again, these values are reported on the log scale, so would need to be exponentiated in order to interpret them on the odds scale. We will do that in a later analysis. The least-squares means for MIX are presented next. When we looked at the overall affect of MIX earlier, it did not have a significant affect on count. However, when we look at the differences in the least-squares means for mixed, almost all of the comparisons are significantly different. This is because these tests did not take into account the overdispersion. The next analysis we'll do will take into account the overdispersion.

mix Least Squares Means

mix	Estimate	Standard Error	DF	t Value	Pr > t
1	3.2716	0.08608	27.45	38.01	<.0001
2	3.1716	0.08726	29	36.35	<.0001
3	3.0641	0.08708	28.76	35.19	<.0001
4	2.8907	0.08955	32.16	32.28	<.0001

Differences of mix Least Squares Means

mix	_mix	Estimate	Standard Error	DF	t Value	Pr > t
1	2	0.09999	0.05602	84	1.79	0.0779
1	3	0.2075	0.05574	84	3.72	0.0004

1	4	0.3809	0.05951	84	6.40	<.0001
2	3	0.1075	0.05755	84	1.87	0.0653
2	4	0.2809	0.06122	84	4.59	<.0001
3	4	0.1734	0.06096	84	2.84	0.0056

Accounting for overdispersion

The following PROC GLIMMIX code accounts for overdispersion.

```
proc glimmix data=spd;
class trt blk mix;
model count=trt mix trt*mix/dist=poisson link=log ddfm=satterth;
lsmeans trt mix/diff;
ods output diffs=diff;
random blk blk*trt;
random _residual_;
run;
```

The code is the same as the previous model, except that we now have an additional random statement. This fits a polynomial Poisson regression model with overdispersion where the variance function $v(\mu)$ is replaced by $\phi v(\mu)$. There is also an ODS output statement. This is for outputting the differences from the LSMEANS in order to calculate the confidence intervals. The code for this will be presented later.

The model and class information are the same as for the previous model. However the output on the dimensions and optimization have changed.

Dimensions	
G-side Cov. Parameters	2
R-side Cov. Parameters	1
Columns in X	40
Columns in Z	32
Subjects (Blocks in V)	1
Max Obs per Subject	112

The dimension has the addition of the R-side covariance parameter, which is the overdispersion parameter. The optimization information has the additional line concerning the residual variance.

Optimization Information	
Optimization Technique	Dual Quasi-Newton
Parameters in Optimization	2
Lower Boundaries	2
Upper Boundaries	0
Fixed Effects	Profiled
Residual Variance	Profiled
Starting From	Data

The iteration history shows that more rounds were needed before convergence was reached. The more parameters that are being estimated, the more iterations that are needed for convergence.

Iteration History						
Iteration	Restarts	Subiterations	Objective Function	Change	Max Gradient	
0	0	5	178.13148738	2.00000000	880.3884	
1	0	3	206.97129676	0.36514487	855.9358	
2	0	4	211.65787334	0.08370848	874.6871	
3	0	2	211.74945361	0.00518381	876.5665	
4	0	2	211.748509	0.00014994	876.6344	
5	0	1	211.74839278	0.00000325	876.6363	
6	0	1	211.74839003	0.00000051	876.6365	
7	0	1	211.7483896	0.00000043	876.6364	
8	0	0	211.74838996	0.00000000	876.6364	

Convergence criterion (PCONV=1.11022E-8) satisfied.

The estimated G matrix is still not positive definite.

Estimated G matrix is not positive definite.

Fit Statistics		
-2 Res Log Pseudo-Likelihood		211.75
Pseudo-AIC (smaller is better)		215.75
Pseudo-AICC (smaller is better)		215.90
Pseudo-BIC (smaller is better)		214.52
Pseudo-CAIC (smaller is better)		216.52
Pseudo-HQIC (smaller is better)		213.05
Pearson Chi-Square		805.57
Pearson Chi-Square / DF		9.59

Note that the Pearson Chi-Square/DF statistic in this model is larger than for the model without accounting for overdispersion. In the model without accounting for the overdispersion, a portion of it went into the BLK*TRT variance component estimate, so the estimate was smaller. When we look at the variance parameter estimates, we see that the BLK*TRT estimate is smaller than the (0.1652) that was estimated in the earlier model. The “Residual (VC)” parameter is the overdispersion parameter estimate.

Covariance Parameter Estimates		
Cov Parm	Estimate	Standard Error
blk	0	.
trt*blk	0.04877	0.04729
Residual (VC)	9.5901	1.7028

Next are the Type III tests of fixed effects. The TRT effect probability in the previous model was 0.03. Remember that the TRT effect is tested over the TRT*BLK variance component. The decrease in the size of that variance component estimate allows for the greater precision.

Type III Tests of Fixed Effects

Effect	Num DF	Den DF	F Value	Pr > F
trt	6	22.61	3.80	0.0092
mix	3	63.44	1.55	0.2115
trt*mix	18	63.44	1.09	0.3802

The treatment least-squares means and differences between least-squares means are presented next. These results are fairly similar to the results that we got with the model without overdispersion. Those treatments that were significantly different before are still different, and those that were not are still not.

trt Least Squares Means

trt	Estimate	Standard Error	DF	t Value	Pr > t
1	2.5062	0.2538	55.67	9.87	<.0001
2	2.9855	0.2089	30.04	14.29	<.0001
3	2.8682	0.2155	33.9	13.31	<.0001
4	3.1401	0.1974	24.48	15.90	<.0001
5	3.4728	0.1766	15.64	19.66	<.0001
6	3.3502	0.1862	19.16	17.99	<.0001
7	3.7016	0.1672	12.52	22.14	<.0001

Differences of trt Least Squares Means

trt	_trt	Estimate	Standard Error	DF	t Value	Pr > t
1	2	-0.4793	0.3287	43.43	-1.46	0.1520
1	3	-0.3620	0.3330	45.33	-1.09	0.2827
1	4	-0.6339	0.3216	40.67	-1.97	0.0555
1	5	-0.9665	0.3092	35.72	-3.13	0.0035
1	6	-0.8440	0.3148	37.73	-2.68	0.0108
1	7	-1.1954	0.3040	33.74	-3.93	0.0004
2	3	0.1173	0.3001	31.96	0.39	0.6986
2	4	-0.1546	0.2874	27.24	-0.54	0.5951
2	5	-0.4873	0.2735	22.5	-1.78	0.0884
2	6	-0.3647	0.2798	24.43	-1.30	0.2046
2	7	-0.7161	0.2676	20.68	-2.68	0.0143
3	4	-0.2719	0.2923	29.13	-0.93	0.3599
3	5	-0.6045	0.2786	24.3	-2.17	0.0400
3	6	-0.4820	0.2848	26.27	-1.69	0.1024
3	7	-0.8333	0.2728	22.43	-3.06	0.0057
4	5	-0.3327	0.2649	19.89	-1.26	0.2237
4	6	-0.2102	0.2714	21.76	-0.77	0.4470
4	7	-0.5615	0.2587	18.13	-2.17	0.0435
5	6	0.1225	0.2566	17.37	0.48	0.6390
5	7	-0.2288	0.2432	14.04	-0.94	0.3627
6	7	-0.3513	0.2502	15.71	-1.40	0.1798

When we look at the MIX results, they are different from the results from the previous model. The estimates for the differences are the same. However the standard error estimates are larger. The probabilities for the differences are in line with the overall MIX effect probability.

mix Least Squares Means

mix	Estimate	Standard Error	DF	t Value	Pr > t
1	3.3185	0.1258	81.35	26.39	<.0001
2	3.2185	0.1334	83.21	24.13	<.0001
3	3.1110	0.1322	83.02	23.53	<.0001

4	2.9376	0.1472	83.96	19.96	<.0001
---	--------	--------	-------	-------	--------

Differences of mix Least Squares Means

mix	_mix	Estimate	Standard Error	DF	t Value	Pr > t
1	2	0.09999	0.1735	63.44	0.58	0.5664
1	3	0.2075	0.1726	63.44	1.20	0.2338
1	4	0.3809	0.1843	63.44	2.07	0.0429
2	3	0.1075	0.1782	63.44	0.60	0.5485
2	4	0.2809	0.1896	63.44	1.48	0.1434
3	4	0.1734	0.1888	63.44	0.92	0.3619

Ignoring the Poisson distribution and treating the data as normally distributed

Before we had software available for analyzing generalized linear mixed models, many researchers treated the data as having an underlying normal distribution and relied on the robustness of the normal distribution. The following code analyzes the data in this way.

```
proc mixed data=spd;
class trt blk mix;
model count=trt mix trt*mix/ddfm=satterth;
lsmeans trt mix/diff;
random blk blk*trt;
run;
```

Only some of the results will be presented. Just as with the GLMM, the BLK variance parameter was so small that the estimate was set to zero.

Covariance Parameter Estimates

Cov Parm	Estimate
blk	0
trt*blk	18.6070
Residual	315.49

The Type III treatment effects are fairly similar to the GLMM.

Type 3 Tests of Fixed Effects

Effect	Num DF	Den DF	F Value	Pr > F
trt	6	21	4.33	0.0054
mix	3	63	2.44	0.0729
trt*mix	18	63	1.10	0.3763

When we look at the differences in the TRT and MIX LSMeans, they are also similar in the probabilities.

Differences of Least Squares Means

Effect	trt	mix	_trt	_mix	Estimate	Standard Error	DF	t Value	Pr > t
trt	1		2		-7.6875	6.9814	21	-1.10	0.2833
trt	1		3		-4.6250	6.9814	21	-0.66	0.5149
trt	1		4		-10.8750	6.9814	21	-1.56	0.1342
trt	1		5		-20.0000	6.9814	21	-2.86	0.0093
trt	1		6		-18.2500	6.9814	21	-2.61	0.0162
trt	1		7		-29.8750	6.9814	21	-4.28	0.0003
trt	2		3		3.0625	6.9814	21	0.44	0.6654
trt	2		4		-3.1875	6.9814	21	-0.46	0.6527
trt	2		5		-12.3125	6.9814	21	-1.76	0.0923
trt	2		6		-10.5625	6.9814	21	-1.51	0.1452
trt	2		7		-22.1875	6.9814	21	-3.18	0.0045
trt	3		4		-6.2500	6.9814	21	-0.90	0.3808
trt	3		5		-15.3750	6.9814	21	-2.20	0.0390
trt	3		6		-13.6250	6.9814	21	-1.95	0.0644
trt	3		7		-25.2500	6.9814	21	-3.62	0.0016
trt	4		5		-9.1250	6.9814	21	-1.31	0.2053
trt	4		6		-7.3750	6.9814	21	-1.06	0.3028
trt	4		7		-19.0000	6.9814	21	-2.72	0.0128
trt	5		6		1.7500	6.9814	21	0.25	0.8045
trt	5		7		-9.8750	6.9814	21	-1.4	0.1719
trt	6		7		-11.6250	6.9814	21	-1.67	0.1107
mix		1		2	2.4643	4.7471	63	0.52	0.6055
mix		1		3	8.3571	4.7471	63	1.76	0.0832
mix		1		4	11.3929	4.7471	63	2.40	0.0194
mix		2		3	5.8929	4.7471	63	1.24	0.2191
mix		2		4	8.9286	4.7471	63	1.88	0.0646
mix		3		4	3.0357	4.7471	63	0.64	0.5248

Going back to the PROC GLIMMIX model accounting for overdispersion, the following code calculates the confidence intervals and then exponentiates them to put them on an interpretable scale.

```

data cl_hat;
set diff;
if effect='trt';
tp=(tinv(.975,df));
cl=exp(estimate-tp*stderr);
cu=exp(estimate+tp*stderr);
phat=exp(estimate);

proc print data=cl_hat noobs;
var trt _trt estimate stderr df probt tp phat cl cu;
run;
    
```

trt	_trt	Estimate	StdErr	DF	Probt	tp	phat	cl	cu
1	2	-0.4793	0.3287	43.43	0.1520	2.01611	0.61922	0.31916	1.20139
1	3	-0.3620	0.3330	45.33	0.2827	2.01370	0.69628	0.35610	1.36140
1	4	-0.6339	0.3216	40.67	0.0555	2.02004	0.53054	0.27707	1.01588
1	5	-0.9665	0.3092	35.72	0.0035	2.02865	0.38040	0.20314	0.71233

1	6	-0.8440	0.3148	37.73	0.0108	2.02487	0.42998	0.22731	0.81335
1	7	-1.1954	0.3040	33.74	0.0004	2.03281	0.30260	0.16312	0.56132
2	3	0.1173	0.3001	31.96	0.6986	2.03704	1.12443	0.61013	2.07226
2	4	-0.1546	0.2874	27.24	0.5951	2.05100	0.85677	0.47516	1.54488
2	5	-0.4873	0.2735	22.5	0.0884	2.07120	0.61431	0.34861	1.08254
2	6	-0.3647	0.2798	24.43	0.2046	2.06198	0.69438	0.38996	1.23646
2	7	-0.7161	0.2676	20.68	0.0143	2.08159	0.48867	0.27998	0.85291
3	4	-0.2719	0.2923	29.13	0.3599	2.04483	0.76196	0.41915	1.38513
3	5	-0.6045	0.2786	24.3	0.0400	2.06254	0.54633	0.30752	0.97059
3	6	-0.4820	0.2848	26.27	0.1024	2.05451	0.61754	0.34400	1.10860
3	7	-0.8333	0.2728	22.43	0.0057	2.07155	0.43459	0.24700	0.76467
4	5	-0.3327	0.2649	19.89	0.2237	2.08673	0.71701	0.41252	1.24623
4	6	-0.2102	0.2714	21.76	0.4470	2.07521	0.81046	0.46147	1.42338
4	7	-0.5615	0.2587	18.13	0.0435	2.09981	0.57036	0.33129	0.98197
5	6	0.1225	0.2566	17.37	0.6390	2.10642	1.13034	0.65832	1.94078
5	7	-0.2288	0.2432	14.04	0.3627	2.14416	0.79548	0.47223	1.33999
6	7	-0.3513	0.2502	15.71	0.1798	2.12311	0.70375	0.41369	1.19719